

ANALYSIS OF MARKET INTEGRATION AND PRICE VARIATION IN MAIZE MARKETING IN OSUN STATE, NIGERIA

ALUFOHAI, G.O.^{1*} and A.A. AYANTOYINBO²

(* corresponding author. Telephone number: 234-805633580, email: ax4grace@yahoo.com)

¹ Department of Agricultural Economics & Extension Services, Faculty of Agriculture, University of Benin, P.M.B. 1154, Benin City, Edo State, Nigeria.

² Department of Agricultural Economics & Extension Services, Faculty of Agriculture, University of Benin, P.M.B. 1154, Benin City, Edo State, Nigeria.

ABSTRACT:

Aims:

The **aim of the** study **was to** examine market integration and price variation in ~~local rice and~~ white maize marketing in Osun State, Nigeria. Specific issues addressed in the study **include** were the determination of existence of co-integration between the rural and urban markets and the leading market between the rural and urban markets for ~~the food grains~~ maize.

Study design: Secondary data **were** used in the is study. The data were sourced from Osun State Development Programme. The average monthly retail price (N/Kg) of local rice covering the period, January, 2000 to December, 2010 (11years) for Osogbo (urban market), Telemu (rural market) and Erin Ijesa (rural market) were used.

Place and Duration of Study: Oson State, South-West Nigeria, between March and July 2011.

Methodology: Coefficient of variation and price correlation coefficient were used to examine the behavior of ~~local rice and~~ white maize market price in urban and rural areas of Osun State. The Augmented Dickey-Fuller (ADF) test was used to investigate stationarity in the pairs of prices while the Johansen co-integration technique was used to determine the existence of co-integration between the markets.

Results: Augmented Dicker Fuller procedure (ADF) indicated that all the variables the individual price series for rice were not stationary at their respective levels, but stationary at their first difference. Both the trace and maximum eigen value statistics indicated the existence of co-integration relation at 5% significant level for the pairs of product prices, implying that ~~rice and~~ maize markets during the study period were linked together and therefore the long-run equilibrium was stable. The results also indicated that urban rice and maize markets did not granger-cause rural rice ~~and maize~~ markets ($p > 0.05$), while rural ~~rice and~~ maize markets granger caused urban ~~rice and~~ maize markets ($p < 0.05$) at 5% level of significance respectively. The error correction model showed significant causality link between the rural and urban markets, suggesting a clear trend in price leadership. It follows that there could be efficiency in transmission of price information among operators if relative stability is attained in the rural markets of ~~local rice and~~ white maize in Nigeria.

Conclusion: The study **concluded** that maize markets in both urban and rural areas of Osun were co-integrated and had short-run and long-run relationships, with rural maize markets occupying the leadership position in price formation and transmission.

Keywords: Causality effect, Co-integration, Granger-caused, Price series, maize.

INTRODUCTION

THE PLACE OF AGRICULTURE IN ENHANCING THE WELLBEING OF THE PEOPLE CANNOT BE OVEREMPHASIZED. THIS IS TRUE BECAUSE, IT IS THE ONLY SECTOR PROVIDING STAPLE FOOD FOR THE INCREASING POPULATION (1). THE NIGERIAN GOVERNMENT'S MAIN FOCUS AT PRESENT IS ON RAISING AGRICULTURAL PRODUCTIVITY, BOTH IN STAPLES FOR LOCAL AND REGIONAL CONSUMPTION, AND FOR A WIDE RANGE OF PRODUCTS. WITH THE CURRENT FALL IN OIL PRICES, THE NIGERIAN GOVERNMENT IS INCREASINGLY GIVING AGRICULTURE A HIGHER PRIORITY IN ITS EFFORT TO REDUCE POVERTY, DIVERSIFY THE NATIONAL ECONOMY AWAY FROM OIL AND TO ENSURE FOOD SECURITY IN ORDER TO IMPROVE THE PER CAPITAL INCOME IN NIGERIA (2)

THE MAJOR FOOD GRAINS IN NIGERIA ARE RICE, MAIZE, SORGHUM, WHEAT, PEARL, MILLET, AND COWPEA WITH RICE RANKING AS THE SIXTH MAJOR CROP IN TERMS OF THE LAND AREA WHILE SORGHUM ACCOUNTEDS FOR 50 PERCENT OF THE TOTAL GRAIN PRODUCTION AND OCCUPIES ABOUT 45 PERCENT OF THE TOTAL LAND AREA DEVOTED TO CEREAL PRODUCTION IN NIGERIA (~~NATIONAL—EXTENSION AGRICULTURAL RESEARCH AND LIAISON STATION,~~ 3). THE RATE OF GROWTH OF NIGERIA'S FOOD PRODUCTION IS 2.5 PERCENT PER ANNUM IN RECENT YEARS, WHILE FOOD DEMAND HAS BEEN GROWING AT THE RATE OF MORE THAN 3.5 PERCENT PER ANNUM DUE TO HIGH RATE OF POPULATION GROWTH (4). FOOD GRAIN MARKETING WARRANTS SPECIAL ATTENTION BECAUSE FOOD GRAINS PLAY AN IMPORTANT ROLE IN NIGERIA AS STAPLES IN MANY HOMES. SECONDLY, THE STRUCTURE AND THE CONDUCT OF THE FOOD GRAIN MARKET CAN AFFECT THE ECONOMY OF THE PEOPLE IN THE COMMUNITIES AND THE NATION'S ECONOMY IN SIGNIFICANT WAYS (5). THE GOVERNMENT OF NIGERIA, REALIZING THE IMPORTANCE OF THE GRAIN SUBSECTOR HAS CONTINUED TO INTERVENE IN STABILIZING THE GRAIN SUBSECTOR THROUGH AGRICULTURAL POLICY REFORMATION. ADETUNJI AND ADESIYAN (6) STATED THAT AGRICULTURAL PRODUCTION PROBLEMS CAN BE OVERCOME THROUGH INTRODUCING NEW TECHNOLOGY AND EFFICIENT MARKETING SYSTEMS, IT IS HOWEVER OBVIOUS THAT INCREASED PRODUCTION WITHOUT CORRESPONDING WELL-DEVELOPED AND EFFICIENT MARKETING SYSTEM MAY AMOUNT TO WASTAGE OF RESOURCES. ~~OLUKOSI (2007), STATED THAT~~ A GOOD AND EFFICIENT MARKETING SYSTEM PROMOTES THE PACE OF ECONOMIC DEVELOPMENT BY ENCOURAGING SPECIALIZATION, WHICH LEADS TO MORE OUTPUT (7). THIS STUDY THEREFORE SOUGHT TO KNOW THE TREND IN PRICE OF ~~SOME SELECTED GRAIN CROPS (RICE AND MAIZE)~~ IN RURAL AND URBAN MARKETS AS WELL AS THE LEVEL OF INTEGRATION BETWEEN ~~THE MARKET~~S FOR ~~THESE SELECTED GRAIN CROPS~~ RICE IN OSUN STATE. THE RESEARCH QUESTIONS INCLUDE (A) HOW DO PRICES RELATE BETWEEN MARKETS? (B) ARE THERE LINKAGES BETWEEN MARKETS IN URBAN AND RURAL AREAS? (C) WHAT IS THE DEGREE OF MARKET INTEGRATION AMONG THESE RURAL AND URBAN MARKETS?

OBJECTIVES OF THE STUDY

THE OVERALL OBJECTIVE OF THE STUDY WAS TO EXAMINE THE EXTENT AND DEGREE OF MARKET INTEGRATION AND PRICE VARIATIONS IN THE OSUN STATE MAIZE ~~AND RICE~~ MARKETS. THE SPECIFIC OBJECTIVES WERE TO:

1. EXAMINE THE PRICE BEHAVIOUR BETWEEN RURAL AND URBAN MARKETS FOR ~~RICE AND~~ MAIZE

2. DETERMINE IF CO-INTEGRATION (LONG RUN EQUILIBRIUM) RELATIONSHIP EXIST BETWEEN ~~LOCAL RICE AND~~ WHITE MAIZE PRICES IN RURAL AND URBAN MARKETS
3. DETERMINE THE LEADING MARKETS BETWEEN URBAN AND RURAL MARKETS
4. DETERMINE THE DEGREE OF MARKET INTEGRATION.

HYPOTHESIS OF THE STUDY

H₀: PRICES OF ~~RICE AND~~ MAIZE IN RURAL MARKETS DO NOT DETERMINE PRICES OF ~~RICE AND~~ MAIZE IN URBAN MARKETS.

2. METHODOLOGY

STUDY AREA

THIS STUDY WAS CARRIED OUT IN OSUN STATE, NIGERIA. OSUN STATE HAS AN ESTIMATED POPULATION OF 3,423,535 (8). ~~THE OSUN STATE IS MADE UP OF 30 LOCAL GOVERNMENT COUNCILS. ITS COORDINATES ARE: 7°30'N 4°30'E 7.500°N 4.500°E, WITH A LAND MASS OF 9,251 KM² (3,572 SQ MI) (9). AND LIES BETWEEN LONGITUDES 40° AND 60° EAST OF THE GREENWICH MERIDIAN AND LATITUDES 50° AND 80° NORTH OF THE EQUATOR.~~ THIS MEANS THAT THE STATE LIES ENTIRELY IN THE TROPICS. THE STATE IS BOUNDED IN THE WEST BY OYO STATE, IN THE NORTH BY KWARA STATE, IN THE EAST BY ONDO AND EKITI STATES AND IN THE SOUTH BY OGUN STATE. AGRICULTURE IS THE TRADITIONAL OCCUPATION OF THE PEOPLE OF OSUN STATE. THE TROPICAL NATURE OF THE CLIMATE FAVOURS THE GROWTH OF A VARIETY OF FOOD AND CASH CROPS. THE MAIN CASH CROPS INCLUDE COCOA, PALM PRODUCE, KOLA, WHILE FOOD CROPS INCLUDE YAM, MAIZE, CASSAVA, MILLET, RICE AND PLANTAIN. THE VEGETATION CONSISTS OF HIGH FOREST AND DERIVED SAVANNAH TOWARDS THE NORTH. THE CLIMATE IS TROPICAL WITH TWO DISTINCT SEASONS. USUALLY THE WET SEASON LAST BETWEEN MARCH AND OCTOBER, WHILE THE DRY SEASON COMES BETWEEN NOVEMBER AND FEBRUARY. MEAN ANNUAL RAINFALL IS BETWEEN 2,000 AND 22,000. MAXIMUM TEMPERATURE IS 32.5°C WHILE THE RELATIVE HUMIDITY IS 79.90 PERCENT (9). THE CHOICE OF OSUN STATE FOR THE STUDY WAS DEEMED TO BE APPROPRIATE BECAUSE OF ITS ANTECEDENT IN AGRICULTURE AND FOOD MARKETING. ~~ACCORDING TO OGUNBODEDE AND OLAKOJO~~ OSUN AND OYO STATES PRODUCE 50% OF THE TOTAL QUANTITY OF MAIZE PRODUCED IN SOUTHWESTERN NIGERIA AND DUE TO THE INFLUX OF FOOD PRODUCED FROM NEIGHBORING STATES OF OYO, KWARA, EKITI, AND ONDO (10). IN THE STUDY, EFFORT WAS MADE TO ANALYZE PRICE TRENDS OF ~~LOCAL RICE AND~~ WHITE MAIZE IN THE RURAL SUPPLY MARKET AND URBAN DEMAND MARKET IN OSUN STATE, NIGERIA WITH THE VIEW TO DETERMINING IF LINKAGES EXISTED BETWEEN THEM AND ASCERTAINING THE NATURE OF THEIR INTERRELATEDNESS.

SOURCES OF DATA

THE SECONDARY DATA USED IN THE STUDY WERE SOURCED FROM OSUN STATE DEVELOPMENT PROGRAMME. THE AVERAGE MONTHLY RETAIL PRICE (₦/KG) OF ~~LOCAL RICE AND~~ WHITE MAIZE COVERING JANUARY, 2000 TO DECEMBER, 2010 (11YEARS) FOR OSOGBO (URBAN MARKET), TELEMU (RURAL MARKET) AND ERIN IJESA (RURAL MARKET) WERE USED.

DATA ANALYSIS: DATA WERE ANALYZED USING E VIEWS SOFTWARE AND STATISTICAL PROCESSES WERE EMPLOYED IN ORDER TO ACHIEVE AN APPROPRIATE ANALYSIS. TO DEAL WITH NATIONAL CURRENCY FLUCTUATIONS, WHICH MAY CAUSE PRICE TO LOOK AS THOUGH THEY ARE INTEGRATED; ALL PRICES WERE QUOTED IN NAIRA PER KILOGRAM (N/KG) AND SERIES OF PRICES WERE ALL DEFLATED BY USING CONSUMER PRICE INDEX (CPI). THE REAL PRICES OBTAINED WERE THEN USED FOR THE ANALYSES.

$$\text{REAL PRICE} = \frac{\text{NOMINAL PRICE} \times 100}{\text{CPI}}$$

DATA COLLECTED WERE ANALYZED USING THE FOLLOWING: COEFFICIENT OF VARIATION AND PEARSON PRICE CORRELATION COEFFICIENT WERE USED TO EXAMINE THE PRICE BEHAVIOR OF RURAL AND URBAN MARKETS OF ~~LOCAL RICE AND~~ WHITE MAIZE IN STUDY AREA. THIS WAS ADAPTED FROM (11, 12) AND COMPUTED AS SHOWN BELOW:

$$\text{PRICE CORRELATION COEFFICIENT (R)} \quad r = \frac{\sum_{t=1}^T (P_{it} - \bar{P}_{it})(P_{jt} - \bar{P}_{jt})}{\sqrt{\sum_{t=1}^T (P_{it} - \bar{P}_{it})^2 \sum_{t=1}^T (P_{jt} - \bar{P}_{jt})^2}}$$

.....(1)

WHERE,

P_{it} AND P_{jt} = PRICE VARIABLES AT TIME T

$\bar{P}_{it}$ AND $\bar{P}_{jt}$ = MEAN OF THE VARIABLES

PRICE CORRELATION COEFFICIENTS WERE USED TO EXAMINE THE STRENGTH OF PRICE LINKAGES ACROSS MARKETS. GIVEN PRICE SERIES FROM TWO MARKETS AT TIME T, P_{it} AND P_{jt} , THE DEGREE OF LINEAR ASSOCIATION BETWEEN THE MARKETS CAN BE MEASURED BY THE SIGN AND MAGNITUDE OF THE CORRELATION COEFFICIENT, R. THE T - STATISTIC WAS USED TO ASCERTAIN IF THE COEFFICIENT BETWEEN PRICES IN THE MARKETS WAS STATISTICALLY DIFFERENT FROM ZERO.

CO EFFICIENT OF VARIATION (CV)

$$\text{CV} = \frac{\text{STANDARD DEVIATION}}{\text{MEAN}} \times 100\% \quad \text{..... (2)}$$

JOHANSEN MULTIVARIATE CO-INTEGRATION PROCEDURE

JOHANSEN DEVELOPED A MULTIVARIATE CO-INTEGRATION METHOD IN 1988, WHICH IS STILL THE MOST SUITABLE APPROACH TO TEST PRICES OF FOOD MARKETS THAT ARE USUALLY ENDOGENOUS AND SIMULTANEOUSLY DETERMINED (13). THE FOUR IMPORTANT POINTS TO BE CONSIDERED BEFORE PERFORMING CO-INTEGRATION TESTS, ACCORDING TO ENDERS (14), ARE:

FIRST, CO-INTEGRATION REFERS TO ONE OR MORE LINEAR COMBINATIONS OF NON-STATIONARY VARIABLES. SECOND, ALL VARIABLES MUST BE INTEGRATED OF THE SAME ORDER. HOWEVER, THIS CONDITION IS NOT NECESSARILY REQUIRED IN ALL CASES. IT IS POSSIBLE THAT VARIABLES ARE INTEGRATED OF DIFFERENT ORDERS. THIRD, THERE MAY BE AS MANY AS N-1 LINEARLY INDEPENDENT CO-INTEGRATING VECTORS IF A LINEAR COMBINATION OF NON-STATIONARY VARIABLES HAS N VARIABLES. THE NUMBER OF CO-INTEGRATING VECTORS IS CALLED THE CO-INTEGRATING RANK (R). IF MORE THAN TWO TIME SERIES ARE CONSIDERED, IT IS POSSIBLE TO HAVE MORE THAN ONE CO-INTEGRATING RANK. FINALLY, CONSIDER

THE CASE IN WHICH EACH VARIABLE CONTAINS A SINGLE UNIT ROOT. BEFORE CONDUCTING THE CO-INTEGRATION TESTS, THE LAG LENGTHS ARE DETERMINED BY USING THE MINIMUM VALUE OF THE AKAIKE INFORMATION CRITERION.

JOHANSEN CO-INTEGRATION PROCEDURE WAS USED TO DETERMINE THE EXISTENCE OF CO-INTEGRATION (LONG RUN EQUILIBRIUM) RELATIONSHIP BETWEEN ~~LOCAL RICE AND~~ WHITE MAIZE PRICES IN THE RURAL AND URBAN MARKETS OF THE STUDY AREA AS FOLLOWS:

TESTING FOR STATIONARITY:

STATIONARY TIME SERIES DATA HAPPEN IF THE AVERAGE, VARIANCE AND COVARIANCE AT ANY LAG ARE STILL CONSTANT AT ANY TIME (15). THE INDIVIDUAL PRICE SERIES ARE TESTED FOR THE ORDER OF INTEGRATION TO DETERMINE WHETHER OR NOT THEY ARE STATIONARY. A NUMBER OF TESTS FOR STATIONARITY ARE AVAILABLE IN THE LITERATURE; THESE INCLUDE THE DICKEY-FULLER (DF) TEST (16), THE AUGMENTED DICKEY-FULLER (ADF) TEST (17) AND THE PHILIPS-PERRON (PP) TEST. A STANDARD TEST FOR NON-STATIONARITY IS THE AUGMENTED DICKEY FULLER (ADF) TEST (17).

FOR EACH PRICE SERIES X_T THE TEST STATISTIC WAS MEASURED BY THE FOLLOWING REGRESSION.

$$\Delta X_T = A + \Delta X_{T-1} + \sum_{k=1}^p B \Delta X_{T-k} + E_T \quad \dots\dots\dots (3)$$

WHERE

- X_T = PRICE AT TIME T
- Δ = FIRST DIFFERENCE OPERATOR.
- T = TIME INDICATOR.
- E_T = THE ERROR TERM.
- Δ, A AND B = PARAMETERS TO BE ESTIMATED.
- K = NUMBER OF LAG OF THE PRICE VARIABLES TO BE INCLUDED.

THE FIRST STAGE IS TO TEST WHETHER EACH SERIES IS STATIONARY I.E. I (0). IF THE NULL HYPOTHESIS OF NON STATIONARITY CANNOT BE REJECTED, THAT IS, THE ABSOLUTE VALUE OF THE ADF STATISTIC IS SMALLER THAN THE CRITICAL AUGMENTED DICKEY FULLER (ADF) VALUE, THEN THE NEXT STAGE IS TO TEST WHETHER THE FIRST DIFFERENCES ARE STATIONARY. IF THE NULL HYPOTHESIS OF NON-STATIONARITY CANNOT BE REJECTED, THEN THE SERIES IS STILL NOT STATIONARY. THEREFORE, DIFFERENCING CONTINUES UNTIL THE SERIES BECOMES STATIONARY AND ORDER NOTED. THE PROCESS IS CONSIDERED STATIONARY IF $\delta / \delta < 1$, THUS TESTING FOR STATIONARITY IS EQUIVALENT WITH TESTING FOR UNIT ROOTS ($\partial < 1$) UNDER THE FOLLOWING HYPOTHESES:

$H_0: \partial = 0$ THE PRICE SERIES IS NON-STATIONARY OR THERE IS EXISTENCE OF UNIT ROOT.

$H_1: \partial \neq 0$ THE PRICE SERIES IS STATIONARY OR THERE IS WHITE NOISE IN THE SERIES.

THE HYPOTHESIS OF NON-STATIONARITY WILL BE ACCEPTED AT 0.01 OR 0.05 LEVELS IF ADF IS GREATER THAN THE CRITICAL VALUE.
SELECTION OF LAG LENGTH:

FOR THE DETERMINATION OF THE LAG LENGTH TO BE INCLUDED IN VAR (VECTOR AUTO REGRESSION) MODEL. AKAIKE'S (1974) (17, 18 AND 19), INFORMATION

CRITERION (AIC), SCHWARZ'S BAYESIAN INFORMATION CRITERION (SBIC), AND HANNAN-QUINN CRITERION (HQIC) ARE USED FOR VAR MODELS. WHEN USING AIC, SBC OR HQIC BASED ON THE ESTIMATED STANDARD ERRORS IN RESPECTIVE EQUATIONS, THE MODEL WITH THE LOWEST VALUE FROM THE AIC, SBC OR HQIC WILL BE CHOSEN.

$$\begin{aligned}\text{SBIC} &= \text{LN}(\Sigma^2) + K/T \text{ LN } T \\ \text{HQIC} &= \text{LN}(\Sigma^2) + 2K/T \text{ LN } T \\ \text{AIC} &= \text{LN}(\Sigma^2) + 2K/T\end{aligned}$$

IN THIS PAPER, THE AIC WAS USED BECAUSE IT HAS THE LOWEST ESTIMATED STANDARD ERROR WHEN COMPARED WITH OTHERS. AIC CAN BE DESCRIBED BY THE FOLLOWING EQUATION:

$$\text{AIC} = \text{LN}(\Sigma^2) + 2K/T \dots\dots\dots (4)$$

WHERE;

$$\begin{aligned}\Sigma^2 &= \text{THE VARIANCE OF THE ESTIMATED RESIDUALS.} \\ T &= \text{THE NUMBER OF PARAMETERS} \\ K &= \text{THE SAMPLE SIZE.}\end{aligned}$$

THE MAXIMUM LAG LENGTH BEGINS WITH 3 LAGS AND PROCEEDS DOWN TO THE APPROPRIATE LAG BY EXAMINING THE AIC, HQIC AND SIC INFORMATION CRITERIA. THE NUMBER OF LAGGED DIFFERENCE TERMS TO BE INCLUDED CAN BE CHOSEN BASED ON T-TEST, F-TEST OR THE AKAIKE'S INFORMATION CRITERION (AIC) (19).

TESTING THE NUMBER OF CO-INTEGRATING RELATIONSHIPS:

JOHANSEN ALSO, PROPOSED TWO LIKELIHOOD RATIO TESTS NAMELY, EIGEN VALUE AND TRACE STATISTIC FOR THE DETERMINATION OF R. IT IS A MAXIMUM LIKELIHOOD RATIO TEST INVOLVING A REDUCED RANK REGRESSION BETWEEN TWO VARIABLES, SAY I(1) AND I(0). A TRACE HAS A NULL HYPOTHESIS OF NUMBER OF CO-INTEGRATING VECTORS BEING LESS THAN OR EQUAL TO R, WHILE ALTERNATIVE HYPOTHESIS IS THAT THERE ARE MORE THAN R CO-INTEGRATING VECTORS. ADDITIONALLY, A MAX HAS A NULL OF R CO-INTEGRATING VECTORS AGAINST R+1 CO-INTEGRATING VECTORS. FOR BOTH TESTS, IF THE TEST STATISTICS IS MORE THAN THE CRITICAL VALUE, WE REJECT THE NULL HYPOTHESIS. TESTING IS CONDUCTED AS A SEQUENCE AND UNDER THE NULL, R = 0, 1,..N-1. WHEN R = 0, FAILING TO REJECT H_0 WILL COMPLETE THE TEST. BUT IF THIS IS NOT THE CASE MEANING WHEN H_0 : R = 0 IS NOT REJECTED, THE TEST CONTINUES UNTIL THE NULL IS NO LONGER REJECTED.

(a) THE TRACE STATISTIC IS COMPUTED (20 AND 21) AS:

$$\lambda_{trace} = T \sum_{i=r+1}^n \ln(1 - \hat{\lambda}_i) \dots\dots\dots (6)$$

WHERE;

$$\begin{aligned}\lambda_i &= \text{ESTIMATED EIGEN VALUE (CHARACTERISTIC ROOTS) OBTAINED FROM } \Pi \text{ MATRIX} \\ T &= \text{THE SAMPLE SIZE.} \\ R &= \text{NUMBER OF CO-INTEGRATING VECTORS} \\ N &= \text{NUMBER OF VARIABLES UNDER CONSIDERATIONS.}\end{aligned}$$

(B) THE MAXIMUM EIGEN VALUE STATISTIC COMPUTED AS:

$$\begin{aligned}\text{MAX}(R/R+1) &= -T \text{ LN}(1 - \lambda_{R+1}) \dots\dots\dots (7) \\ T &= \text{THE SAMPLE SIZE}\end{aligned}$$

(Λ_{R+1}) = ESTIMATED EIGEN VALUES (CHARACTERISTIC ROOTS) OBTAINED FROM THE Π MATRIX

H_0 : THERE IS NO CO-INTEGRATING VECTOR BETWEEN THE-M ESTIMATED PRICES FOR MAIZE.

H_A : THERE IS CO-INTEGRATING VECTOR BETWEEN THE-M ESTIMATED PRICES FOR MAIZE.

IF THE VALUE OF Λ TRACE AND Λ MAX EXCEED THE CRITICAL VALUE, REJECT THE NULL HYPOTHESIS AND ACCEPT THE ALTERNATIVE HYPOTHESIS OF MORE CO-INTEGRATION VECTORS AT 0.05 OR 0.01 LEVEL.

ABSENCE OF A CO-INTEGRATING RELATIONSHIP SPOTS NONEXISTENCE OF LONG-RUN RELATIONSHIP.

VECTOR ERROR CORRECTION MODEL (VECM):

IF PRICES ARE INTEGRATED OF THE SAME ORDER AND PRICES OF EACH MODEL ARE CO-INTEGRATED, A VECTOR ERROR CORRECTION MODEL (VECM) IS APPROPRIATE TO DETERMINE THE MULTIVARIATE RELATIONSHIPS AMONG PRICES.

JOHANSEN DEFINED TWO MATRICES A AND B, SUCH THAT $\Pi = AB'$, WHERE BOTH A AND B ARE $(N \times R)$ MATRICES. THE PROCEDURE IS BASED ON MAXIMUM LIKELIHOOD ESTIMATION OF THE ERROR CORRECTION MODEL AND EACH TWO-VARIABLE SYSTEM IS MODELED AS A VECTOR AUTO REGRESSION (VAR) AS IN THE FOLLOWING EQUATION (22, 23, 24):

$$X_T = \mu + \sum_{i=1}^p \Gamma_i \Delta X_{T-i} + \Pi X_{T-K} + E_T + B_T \dots \dots \dots (8)$$

WHERE;

X = THE VECTOR OF ENDOGENOUS VARIABLES

Γ_i = THE MATRIX OF SHORT RUN COEFFICIENTS

Π = THE MATRIX OF LONG-RUN COEFFICIENTS

E_T = THE VECTOR OF INDEPENDENT AND NORMALLY DISTRIBUTED ERRORS.

K = NUMBER OF LAGS, AND SHOULD BE ADEQUATELY LARGE ENOUGH BOTH TO CAPTURE THE SHORT-RUN DYNAMICS OF THE UNDERLYING VAR AND TO PRODUCE NORMALLY DISTRIBUTED WHITE NOISE RESIDUALS.

IF THE COEFFICIENT MATRIX Π HAS REDUCED RANK $R < N$, THEN THERE EXIST $N \times R$ MATRICES A AND B EACH WITH RANK R SUCH THAT $\Pi = AB'$ AND $B'X_T$ IS STATIONARY R IS THE NUMBER OF CO-INTEGRATING RELATIONSHIPS, THE ELEMENTS OF A ARE KNOWN AS THE ADJUSTMENT PARAMETERS IN THE VECTOR ERROR CORRECTION MODEL AND EACH COLUMN OF B IS A CO-INTEGRATING VECTOR. IT CAN BE SHOWN THAT FOR A GIVEN R , THE MAXIMUM LIKELIHOOD ESTIMATOR OF B DEFINES THE COMBINATION OF X_{T-1} THAT YIELDS THE R LARGEST CANONICAL CORRELATIONS OF ΔX_T WITH X_{T-1} AFTER CORRECTING FOR LAGGED DIFFERENCES AND DETERMINISTIC VARIABLES WHEN PRESENT. JOHANSEN PROPOSED TWO DIFFERENT LIKELIHOOD RATIO TESTS OF THE SIGNIFICANCE OF THESE CANONICAL CORRELATIONS AND THEREBY THE REDUCED RANK OF THE Π MATRIX.

THE PROCEDURE FOR TESTING CO-INTEGRATION IS BASED ON THE ERROR CORRECTION MODEL (ECM) REPRESENTATION OF X_T GIVEN BY (25; 22, 23, 24):

$$\Delta X_T = \mu + \sum_{i=1}^{p-1} \Gamma_i \Delta X_{T-i} + \Pi X_{T-K} + E_T + B_T \quad \dots\dots\dots (9)$$

WHERE;

Δ = THE DIFFERENCE OPERATOR
 X_T = (NX1) VECTOR OF I (1) (I.E INTEGRATED OF ORDER ONE) PRICES
 Γ_i = - (I - Π_1 - ,K -1)
 i = 1,2,....., K - 1
 Π = - (I - Π_1 - Π_K) EACH OF Π_i IS AN (N X) MATRIX OF PARAMETERS
 K = NUMBER OF LAGS
 E_T = AN IDENTICAL AND INDEPENDENTLY DISTRIBUTED N-
 DIMENSIONAL VECTOR OF
 RESIDUALS WITH ZERO MEAN AND VARIANCE MATRIX
 B = CO-INTEGRATING VECTOR (CONTAINING THE LONG-RUN)
 μ = CONSTANT TERM
 T = TIME TREND.

SINCE X_{T-K} I (1), BUT ΔX_T AND ΔX_{T-1} VARIABLES ARE I(0) (I.E., INTEGRATED OF ORDER ZERO), EQUATION (9) WILL BE BALANCED IF $\Delta X_T - K$ IS I(0). SO, IT IS THE Π MATRIX THAT CONVEYS INFORMATION ABOUT THE LONG-RUN RELATIONSHIP AMONG THE VARIABLES IN X_T . THE RANK OF Π , R, DETERMINES THE NUMBER OF CO-INTEGRATING VECTORS, AS IT DETERMINES HOW MANY LINEAR COMBINATIONS OF X_T ARE STATIONARY. IF $R = N$, THE VARIABLES ARE STATIONARY IN LEVELS. IF $R = 0$, NO LINEAR COMBINATION OF X_T IS STATIONARY.

IF $0 < \text{RANK}(\Pi) = R < N$, AND THERE ARE $N \times R$ MATRICES A AND B SUCH THAT $\Pi = AB'$ THEN IT CAN BE SAID THAT THERE ARE R CO-INTEGRATING RELATIONS AMONG THE ELEMENTS OF X_T . THE CO-INTEGRATING VECTOR B HAS THE PROPERTY THAT $B'X_T$ IS STATIONARY EVEN THOUGH X_T ITSELF IS NON-STATIONARY. THE MATRIX A MEASURES THE STRENGTH OF THE CO-INTEGRATING VECTORS IN THE ECM, AS IT REPRESENTS THE SPEED OF ADJUSTMENT PARAMETER.

EMPIRICAL MODEL:

FOR THIS STUDY, IT WAS HYPOTHESIZED THAT RURAL AND URBAN MARKET PRICES FOR **LOCAL RICE AND** WHITE MAIZE ARE JOINTLY DETERMINED OR ENDOGENOUS, GIVEN AN IMPLICIT REPRESENTATION OF THE MODEL WITH TWO ENDOGENOUS VARIABLES WITHOUT EXOGENOUS VARIABLES AS (23, 24).

$$X_T = (\text{LN_RP}_T, \text{LN_UP}_T) \quad \dots\dots\dots (10)$$

WHERE;

X_T
 LN_RP_T = NATURAL LOG OF RURAL MARKET PRICE
 LN_UP_T = NATURAL LOG OF URBAN MARKET PRICE

FROM EQUATION (10) ABOVE, THE LONG-RUN CO-INTEGRATING EQUATION CAN BE SPECIFIED EXPLICITLY FOR RURAL MARKET PRICE AS;

$$\text{LN_RPT} = \Omega_0 + \Omega_1 \text{LNUP}_T + Y_T \quad \dots\dots\dots (11)$$

WHERE;

Ω_0 = THE LOG OF A PROPORTIONALITY COEFFICIENT, A
 CONSTANT TERM CAPTURING
 THE TRANSPORTATION AND OTHER FORMS OF COST

Ω_1 = LONG RUN COEFFICIENT DEPRECATING THE
 RELATIONSHIP BETWEEN RURAL AND
 URBAN MARKET PRICES
 Y_T = RANDOM ERROR TERM
 IF $\Omega_1=0$ THEN THERE IS NO RELATIONSHIP
 IF $0 < \Omega_1 < 1$ THERE IS A RELATIONSHIP BUT THE RELATIVE PRICE IS
 NOT CONSTANT,
 MEANING THAT THE GOODS WILL BE IMPERFECT
 SUBSTITUTES.
 IF $\Omega_1=1$ THERE IS RELATIONSHIP WITH CONSTANT RELATIVE
 PRICE, MEANING THAT THE
 LAW OF ONE PRICE HOLDS AND GOODS ARE PERFECT
 SUBSTITUTES.

EQUATION (11), DESCRIBES A CASE WHERE PRICES ADJUST IMMEDIATELY. IF
 HOWEVER, A DYNAMIC ADJUSTMENT PATTERN IS EXPECTED IN PRICES, IT WILL BE
 ACCOUNTED FOR BY INTRODUCTION OF LAGS OF THE TWO PRICES, BUT EVEN AT
 THAT, THE LONG-RUN RELATIONSHIP BETWEEN PRICES WILL TAKE THE SAME FORM
 DEPICTED IN EQUATION (11) ABOVE.

VECM MODEL IN THIS STUDY WAS ESTIMATED AS (22, 23, 24):

$$\Delta RP_t = \psi_{10} + \sum_{i=1}^p \psi_{11i} \Delta RP_{t-i} + \sum_{i=1}^p \psi_{12i} \Delta UP_{t-i} - \rho (RP_{t-1} - UP_{t-1}) + Y_{1t} \quad (12)$$

$$\Delta UP_t = \psi_{20} + \sum_{i=1}^p \psi_{21i} \Delta RP_{t-i} + \sum_{i=1}^p \psi_{22i} \Delta UP_{t-i} - \rho (RP_{t-1} - UP_{t-1}) + Y_{2t} \quad (13)$$

WHERE;

Δ = THE DIFFERENCE OPERATOR
 RP AND UP = RURAL AND URBAN MARKETS PRICES
 ψ_{11} AND ψ_{12} = SHORT RUN COEFFICIENTS
 ρ = ERROR CORRECTION INSTRUMENT MEASURING THE
 SPEED OF
 ADJUSTMENT FROM THE SHORT-RUN STATE OF
 DISEQUILIBRIUM TO THE
 LONG-RUN STEADY-STATE EQUILIBRIUM
 Y_T = AN ERROR TERM ASSUMED TO BE DISTRIBUTED AS
 WHITE NOISE
 ψ_{10} AND ψ_{20} = CONSTANTS.

GRANGER CAUSALITY TEST:

THE GRANGER CAUSALITY TEST WAS USED TO DETERMINE THE LEADING MARKETS
 BETWEEN URBAN AND RURAL MARKETS. GRANGER CAUSALITY PROVIDES
 ADDITIONAL EVIDENCE AS TO WHETHER, AND IN WHICH DIRECTION, PRICE
 INTEGRATION AND TRANSMISSION IS OCCURRING BETWEEN TWO PRICE SERIES OR
 MARKET LEVELS. THIS IS BECAUSE ONE GRANGER CAUSAL RELATIONSHIP MUST
 EXIST IN A GROUP OF CO-INTEGRATED SERIES (8). WHEN GRANGER CAUSALITY RUN
 ONE WAY (UNI-DIRECTIONAL), THE MARKET WHICH GRANGER-CAUSES THE OTHER
 IS TAGGED THE EXOGENOUS MARKET. EXOGENEITY CAN BE WEAK OR STRONG.

WEAK EXOGENEITY OCCURS WHEN THE MARGINAL DISTRIBUTION OF $X_{1(T-1)}$ AND $X_{J(T-1)}$ ARE SIGNIFICANT, WHILE STRONG EXOGENEITY OCCURS WHEN THERE IS NO SIGNIFICANT GRANGER-CAUSALITY FROM THE OTHER VARIABLE. IT COULD ALSO BE BI-DIRECTIONAL WHICH INDICATES THAT BOTH SERIES INFLUENCE EACH OTHER (E.G. X CAUSES Y AND Y ALSO CAUSES X).

THE GRANGER MODEL USED IN THIS STUDY IS REPRESENTED BY:

$$RP_t = \alpha_0 + \sum_{i=1}^m \alpha_i UP_{t-i} + \sum_{j=1}^n \beta_j RP_{t-j} + \varepsilon_t \quad 14$$

WHERE;

N = NUMBER OF OBSERVATION

M = NUMBER OF LAG

RP_T = RURAL MARKET PRICE

UP_T = URBAN MARKET PRICE

A AND B = PARAMETERS TO BE ESTIMATED.

H_0 : PRICE OF ~~LOCAL RICE AND~~ WHITE MAIZE IN ONE MARKET DOES NOT DETERMINE (GRANGER CAUSE) THE PRICE IN THE OTHER MARKET

H_A : PRICE OF ~~LOCAL RICE AND~~ WHITE MAIZE IN ONE MARKET DOES DETERMINE THE PRICE IN THE OTHER MARKET (NOT GRANGER CAUSE)

WALD TEST FOR MARKET INTEGRATION:

THE TYPE AND DEGREE OF MARKET INTEGRATION WAS DETERMINED BY THE STATISTICAL SIGNIFICANCE OF THE ESTIMATED PARAMETERS BASED ON THE RESULTS OF THE SET OF HYPOTHESIS USING THE F-STATISTIC OF THE WALD TESTS RESTRICTIONS. THE RESTRICTIONS WERE TESTED ON THE OLS (ORDINARY LEAST SQUARE) ESTIMATION OF THE FOLLOWING EQUATION (8, 23 AND 24):

$$X_t = \eta + \sum_{k=1}^K A_k X_{t-k} + \varepsilon_t$$

WHERE;

T = 1, 2,... REFERS TO THE MONTHS FROM JANUARY 2000 TO DECEMBER, 2010

X_T = N × 1 VECTOR OF THE LOGARITHMIC PRICES AT TIME T ($X_T = X_{1T}, X_{2T}, \dots, X_{NT}$)

A_i = N × N MATRICES OF PARAMETERS;

η = N × 1 VECTOR OF INTERCEPT TERMS

ET = N×1 VECTOR OF ERROR TERMS,

K = THE LAG LENGTH

ET = THE VECTOR OF ERROR TERM

(A) LONG-RUN MARKET INTEGRATION:

H_0 : ~~LOCAL RICE OR~~ WHITE MAIZE MARKET PRICES ARE INTEGRATED IN THE LONG-RUN.

H_A : ~~LOCAL RICE OR~~ WHITE MAIZE MARKET PRICES ARE NOT INTEGRATED IN THE LONG-RUN.

(B) SHORT -RUN MARKET INTEGRATION:

H_0 : A PRICE CHANGE IN A MARKET IS IMMEDIATELY TRANSMITTED TO THE OTHER MARKET.

H_A: A PRICE CHANGE IN A MARKET IS NOT IMMEDIATELY TRANSMITTED TO THE OTHER MARKET.

3. RESULTS AND DISCUSSION

3.1 Price behavior of ~~rice and~~ maize in the study area

Average annual retail prices of ~~local rice and~~ white maize: The behavior of the average monthly retail price of ~~local rice and~~ maize from 2000 to 2010 are presented in Tables 1 and 2. It could be seen that ~~in Osogbo market, prices ranged from an average of about N 40.2/kg in 2000 to a peak of N 99 /kg in 2010. In Erin Ijesa market, prices ranged from N 36.6/kg in 2000 to N 94/kg in 2010 while~~ retail prices of white maize in Telemu market ranged from ~~N~~37/kg in 2000 to ~~N~~ 95/kg in 2010. It was observed, that Erin Ijesa market had the lowest period average price (~~N~~36.6 /kg) while Osogbo market had the highest (~~N~~ 99/kg).

~~Table 1: Average annual retail price of local rice (2000 – 2010) in N/kg.~~

Year	Osogbo (Urban)	Erin Ijesa (Rural)	Telemu (Rural)
2000	48.4	47.2	49.5
2001	58.5	53.4	60.5
2002	62.6	58.9	63.5
2003	72.5	67.2	73.5
2004	85.8	82.5	88.0
2005	95.6	90.0	98.5
2006	115	110.9	117.3
2007	125.4	120	126.5
2008	137.0	132.8	136.2
2009	145.4	139	143.2
2010	160	146.5	162.0
Average	100.5	95.3	101.7
Overall average	99.1		

~~Source: Computed, employing Price data series from OSSADEP.~~

~~Table 2-1: Average annual retail price of white maize 2000 – 2010 in ~~N~~/kg.~~

Year	Osogbo	Erin Ijesa	Telemu
2000	40.2	36.6	37.0
2001	45.2	42.8	42.5

2002	53.5	49.8	48.5
2003	56.4	54.3	55.2
2004	64.1	62.1	62.9
2005	72.7	69.2	68.3
2006	79.5	73.4	74.0
2007	86.2	82.0	81.5
2008	92.3	88.0	85.2
2009	95.6	90.0	90.5
2010	99.0	94.0	95.0
Average	71.3	67.4	67.3
Overall Average	68.7		

Source: Computed employing Price data series from OSSADEP (2012)

The average annual retail price of white maize in urban area was higher than the price in rural area, while in the case of local rice, the price is higher at Telemu market, consumers' market in the rural area than Erin Ijesa a producers' market in the rural area. The difference in mean is expected and in this case statistically significant ($p < 0.01$). Among other things, they could represent the extra cost, including transportation and transactions, incurred by the marketing agents, as well as marketing margins. It has been argued (7), that given the high cost of transactions and the risk to invested capital, the margins of the marketing agents could be considered reasonable. These findings collaborate agree with the findings of Ojiako et al. (2012 23) in their analysis of the spatial integration of cassava product market price in Nigeria, where it was reported that the mean price value of *Lafun* in the urban market was higher than prices in the rural markets.

Variability in average retail prices of local rice and white maize:

The average annual retail price of white maize in urban area was higher than the price in rural area, while in the case of local rice, the price is higher at Telemu market, a consumers' market in the rural area than Erin Ijesa a producers' market in the rural area. Variability in average monthly retail price of white maize is presented in Table 2. Telemu market (a rural market dominated by retailers) had the highest coefficient of variation. The rural markets had a higher coefficient of variation than the urban market. This indicates that the retail price for white maize in the urban market was more stable than what obtained in the rural markets (producer areas). This is because most marketers preferred to sell their produce in the urban area and this could eventually lead to scarcity in the rural area and high price of produce in the rural area, as suggested by (Olukosi et al; oladapo).

Table 32: Variability in average annual retail prices of local rice (2000 - 2010)

Markets	Average Price (₦ /kg)	Coefficient of Variation (%)
Osogbo	100.5	38.2
Erin	95.3	36.0
Telemu	101.7	40.3

Source: Rice price series from January, 2000 to December, 2010 collected from OSSADEP (2012).

Variability in average monthly retail price of white maize is shown in Table 4.3. Monthly prices were also observed to be low during the period covered by the study. Prices varied from 27.9 % percent in Osogbo (a consuming urban area), 28.1 percent in Erin Ijesa (producing rural area) to 28% percent, in Telemu (a producing rural area). There wasn't much variability in average monthly retail price of white maize for the period covered by the study. Prices varied from 27.9 percent in Osogbo (a consuming urban area), 28.1 percent in Erin Ijesa (producing rural area) to 28.0 percent in Telemu (a producing rural area).

Table 4.3: Variability in average annual retail prices of white maize (2000 - 2010)

Markets	Average Price (N /kg)	Coefficient of Variation (%)
Osogbo	71.3	27.9
Erin	67.4	28.1
Telemu	67.3	28.0

Source: Maize price series from January, 2000 to December, 2010 collected from OSSADEP (2012).

Price correlation co-efficient for rice and maize in Osun State: Pearson price correlation analysis was used to determine the behaviour of market price between local rice and white maize in rural and urban markets. The results (Table 5.4) were calculated as 0.90 between Erin and Osogbo, 0.85 between Osogbo and Telemu and 0.35 between Erin and Telemu. Results were significant at 1 percent level except for 0.35 which was significant only at 5 percent level. The results of the analysis of the correlation between market prices of local rice in urban and rural area, is shown in Table 5. The results showed a correlation coefficient of 0.86 between Erin Ijesa and Osogbo market prices, 0.90 between Osogbo and Telemu, and 0.87 between Erin and Telemu. The coefficients were significant at 1 percent level.

Table 5: Pearson correlation between retail prices of local rice in the selected markets

	Osogbo	Erin Ijesa	Telemu
Osogbo	1.00	0.86	0.90
Erin Ijesa	0.86	1.00	0.87
Telemu	0.90	0.87	1.00

Source: Rice price series from January, 2000 to December, 2010 collected from OSSADEP (2012)

Table 3.4: Pearson correlation between retail prices of white maize in selected markets

	Osogbo	Erin Ijesa	Telemu
Osogbo	1	0.92	0.85
Erin	0.92	1	0.35
Telemu	0.85	0.35	1

Source: Maize price series from January, 2000 to December, 2010 collected from OSSADEP (2012).

The high and significant correlation of the price series ~~calculated as 0.90 and 0.86~~ is an indication of co-movement in the prices. The positive correlation showed that an increase in the retail price in one market would follow the price increase in the other market. This could be possible due to the transmission of market information by marketers ~~through various means, particularly~~ via the use of mobile phones, coupled with the short distance between markets. ~~This could also suggest the possible existence of relative price elasticities in these markets.~~ This corroborates the findings of Oladapo and Momoh (22), which noted high correlation coefficients of pineapple prices in Edo, Oyo and Lagos States' rural and urban markets.

JOHANSEN MULTIVARIATE CO-INTEGRATION TEST RESULT

Testing for Stationarity

~~Stationarity test for rice and maize in Osun state: As presented in Table 7, the ADF test statistics calculated at price levels for the local rice price series were -1.82, -1.88 and -1.81 for Osogbo, Erin Ijesa and Telemu respectively. The results indicate that all the variables were not stationary at their level. The null hypothesis of non-stationarity could not be accepted at the probability of 5 % level of significance. Therefore, the null hypothesis of non-stationary was accepted for all the variables at their levels. The values in first differences were -7.66, -6.22 and -8.39 in Osogbo, Erin Ijesa and Telemu respectively. When first-differenced, however, the null hypothesis of non-stationarity was rejected in favour of the alternative as the values of the ADF t-statistics were greater in absolute term than the critical value. This result is necessary and sufficient for a test of co-integration of the price series.~~ As presented in Table 8.5 below, the ADF test statistics calculated at price levels for the rural markets were - 1.42 for Erin and -1.36 for Telemu, while it was -1.729 3 at Osogbo (an urban market). In the first difference, the statistics were calculated as -5.336 4 and -5.923 2 for rural and -6.769 for urban markets. When compared with the ADF test critical value, the result showed that all the series were non-stationary and integrated of order one, that is, they were also I(1) series.

~~Table 7: Unit Root Test on rice price series~~

Market	Price Levels I(0)	First differences I(1)
Osogbo	-1.85	-7.66
Erin Ijesa	-1.81	-6.22
Telemu	-1.88	-8.39
ADF Test Critical values at 5%	-3.89	-3.46
ADF Test Critical Values at 1%	-3.64	-3.64

Source: ~~Rice price series from January, 2000 to December, 2010 collected from OSSADEP (2012).~~

As presented in Table 8-5 below-, ADF test statistics calculated at the price level for the rural markets were - 1.42 for Erin and -1.36 for Telemu, while it was -1. 729 73 at Osogbo (an urban market). In the first difference, the statistics were calculated as -5. 336-35 and - 5. 923- 93 for rural and -6. 760- 76 for the urban markets. When compared with the ADF test critical value, the result showed that all the series were non-stationary and integrated of order one, that is, they were also I(1) series.

Table 8-5: Unit -Root Test on ~~Maize maize Price price Series series~~

Market	Test as Level I(0)	Test at First differences I(1)
Osogbo	-1.172954 3	-6.769
Telemu(Rural)	-1.42244	-5.336 4
Erin Ijesa(Rural)	-1.367	-5.928 3
ADF Test Critical values at 5%	-3.893	-3.469
ADF Test Critical Values at 1%	-4.061	-4.062

Source: ~~Maize price series from January, 2000 to December, 2010 collected from OSSADEP (2012)~~

The results of stationarity test for the ~~rice-and~~ white maize, using Augmented Dickey-Fuller (ADF) unit roots tests indicated that all the variables were not stationary at their levels. The calculated ADF statistic was less than the critical ADF values at both 5% percent and 1% percent levels of significance. Therefore, the null hypothesis of non-stationary was accepted for all the variables at their levels. When first-differenced, however, the null hypothesis of non-stationarity was rejected in favour of the alternative as the calculated ADF values became higher than the critical values at both 5% percent and 1% percent levels. The findings corroborate earlier findings that food commodity price series are mostly stationary of order 1 i.e. I(1) (28, 22, 26 ,27). The result was explained by the fact that most food price series had trends in them because of inflation and therefore exhibited non-stationarity.

Selection of ~~Lag- lag Length- length~~

Akaike Information Criterion (AIC) test (Dickey and Fuller, 1981; Said and Dickey, 1984), suggested that the value K = 1 is the appropriate specification for the order of VAR model. The use of one lag on the model of the economy means implies that all variables in the model influenced each other not only in the present period (from the year 2000 to 2010) but these variables were also interrelated in the period before the year 2000. in-a previous period. These results are in line with the findings of Desi and Yulius (29) which showed that the use of lag 1 is suitable in the co-integration procedure.

Testing for ~~Co-Integration co-integration~~ between ~~Urban urban and Rural rural Market market Prices-prices~~: Both Trace and Maximum Eigen value statistics indicate the existence of co-integrating relationship at 5% percent significant level for ~~maize both products~~. The result of the Johansen's Maximum Likelihood co-integration test is shown in Table 9 6. The result, based on the both the trace test and maximum Eigen value test, showed the existence of two co-integrating vectors and the rejection of the null hypothesis of r = 0. Comparing the trace and Eigen statistic with the corresponding critical values, it can be seen that the null hypothesis of no co-integrating relationship

can be rejected at the 5% percent significance level for the ~~local-rice~~ market prices for white maize.

Market integration lends itself to co-integration interpretation with its presence being evaluated by means of co-integration tests (21). Thus, the result indicates that ~~maize the grain~~ markets in Osun State during the study period were co-integrated, and there existed long-run equilibrium. This finding is supported by the earlier studies carried out by Oladapo and Momoh (22), who concluded that grain market prices within Oyo State are highly co-integrated and the findings of Ojiako et al. (23), that long-run equilibrium existed within the spatial integration of cassava products market in Nigeria.

Table 9: ~~Testing for number of co-integration relations (rice)~~

H ₀ :	H _A :	—5%	Prob**	Hypothesized	
λ Trace tests		λ Trace value	Critical value	No. of CE(s)	
r = 0	r > 0	-70.27	-29.80	-0.00	None*
r ≤ 1	r > 1	-35.37	-15.50	-0.00	At most 1*
r ≤ 2	r > 2	-2.07	-3.84	-0.15	At most 2
λ max tests		λ max value			
r = 0	r = 1	-34.90	-29.80	-0.00	None*
r = 1	r = 2	-33.29	-15.50	-0.00	At most 1*
r = 2	r = 3	-2.07	-3.84	-0.15	At most 2

Source: ~~Price series from January, 2000 to December, 2010; collected from OSSADEP~~

Table 10 6: Testing for number of co-integration relations (maize)

H ₀ :	H _A :		5%	Prob**	Hypothesized
λ Trace tests		λ Trace value	Critical value		No. of CE(s)
r =0	r >0	52.59	29.80	0.00	None*
r ≤1	r >1	20.30	15.50	0.01	At most 1*
r ≤2	r >2	4.80	3.84	0.03	At most 2
λmax tests		λ max value			
r = 0	r =1	32.29	29.80	0.00	None*
r =1	r =2	15.49	15.19	0.01	At most 1*
r =2	r =3	4.80	3.84	0.03	At most 2

Source: Price series from January, 2000 to December, 2010; collected from OSSADEP (2012).

The result of the Johansen's Maximum Likelihood Co-Integration test is presented in Table 11 7. Comparing the trace and Eigen statistics with the corresponding critical values, it can be seen that the null hypothesis of no co-integrating relationship can be rejected at the 5% percent significance level for the ~~local-rice white maize~~ market prices. The results from the trace and Eigen test indicate that there are co-integrating vectors. This implies that there existed a long-run relationship between the variables.

Testing for short-run market integration with a Vector Error Correction Model

The result of the Vector Error Correction Model shows that if there is a positive deviation for the long-run equilibrium, the market tends to respond with a decrease in the rural price or an increase in the urban price. The urban price appears to respond faster than the rural price. The adjustment coefficient is statically significant at 1% percent for urban market price for ~~both local-rice and~~ white maize suggesting that the rural price is weakly exogenous, This implies that movement in the rural price was less affected by

price in the urban market while movement in the urban price was dictated by events in the rural markets. This means that the long-run equilibrium in the **local rice and** white maize market, after an exogenous shock, is restored primarily by corrections made by the urban market prices

Table 11 7: Estimation of the dynamics in the short-run by using VECM for **rice and** maize

Error correction	Δ OSO	Δ ERIN	Δ TEL
Rice			
CointEq1	0.04 (0.18) [-2.12]	-0.18 (0.18) [-1.76]	0.06 (0.18) [-0.31]
Maize			
CointEq1	0.04 (0.21) [-1.90]	-0.13 (0.19) [-1.68]	-0.15 (0.20) [1.76]

Source: Price series from January, 2000 to December, 2010; collected from OSSADEP (2012).

Note: All figures in brackets (...) are standard errors and all figures in parenthesis [...] are t-values.

Granger causality test

The result of the Granger causality analysis (Table 12) indicated that, Osogbo (urban) local rice market price did not determine the Erin Ijesa (rural) market price. Although, Erin Ijesa market price determined the urban market price, while Osogbo local rice market price determined Telemu market price. This result is consistent with the findings of Ojiako et al. (2012) in the study of spatial integration and price transmission in selected cassava products' markets in Nigeria, that Granger causality runs from rural to the urban markets and not the other way round.

TABLE 12: Pair-wise Granger Causality Test for Local Rice Market

Null hypothesis	F-statistics	Probability Value
MERIN does not Granger Cause MOSO	-6.31	0.00***
MOSO does not Granger Cause MERIN	-0.37	0.67
MTEL does not Granger Cause MOSO	-0.33	0.77
MOSO does not Granger Cause MTEL	10.19	0.00***
MTEL does not Granger Cause MERIN	-0.58	0.78
MERIN does not granger cause MTEL	-6.14	0.00***

***Significant at 1% probability level

Source: Rice price series from January, 2000 to December, 2010; collected from OSSADEP

The result of Granger causality test for white maize at 5% significant level is shown in Table 13 8. From the result of the analysis, Osogbo white maize market price did not determine Erin Ijesa and Telemu market price. Although, Erin Ijesa and Telemu market prices determined the Osogbo market price.

Price behavior of **rice and** maize in the study area

Average annual retail prices of local rice and white maize: The behavior of the average monthly retail price of **local rice and** maize from 2000 to 2010 are presented in Tables 1 and 2. It could be seen that in Osogbo market, prices ranged from an average of about N 40.2/kg in 2000 to a peak of N 99 /kg in 2010. In Erin Ijesa market, prices ranged from N 36.6/kg in 2000 to N 94/kg in 2010 while retail prices of white maize in Telemu market ranged from N37/kg in 2000 to N 95/kg in 2010. It was observed, that Erin Ijesa market

had the lowest period average price (N=36.6 /kg) while Osogbo market had the highest (N=99/kg).

Table 13 8: Pair-wise Granger Causality Test for white maize market price

Null Hypothesis	F-Statistics	P-Value
RERIN does not Granger Cause ROSO	4.27	0.02
ROSO does not Granger Cause RERIN	0.18	0.22
RTELE does not Granger Cause ROSO	5.83	0.00
ROSO does not Granger Cause RTELE	0.04	0.02
RTELE does not Granger Cause RERIN	0.18	2.12
RERIN does not granger cause RTELE	0.22	0.54

Source: Maize price series from January, 2000 to December, 2010; collected from OSSADEP (2012).

Wald Test for market integration

white maize show that they are not significant even at 1%. The long-run and short-run null hypotheses that local rice and white maize market prices are integrated and a price change in a market is immediately transmitted to other markets, respectively, therefore cannot be rejected. The results mean that there existed both long-run and short-run market integrations between Osogbo and the other selected markets. Thus, changes in food grain price in rural markets would cause food grain price in urban markets to adjust immediately.

The Wald test restriction of the F-statistic was applied to determine market integration in the local rice ~~local rice and~~ white maize markets (Tables 14 and 15 9). The F statistical values of 0.02 and 0.19 with probability values of 0.86502 and 0.7521 respectively, for local rice whereas it was 0.25 and 0.57 F statistic value with probability values of 0.9573 6 and 0.6210 for white maize which shows that they are not significant even at 1%. The long-run and short-run null hypotheses that ~~local rice and~~ white maize market prices are integrated and a price change in a market is immediately transmitted to other markets, respectively, therefore cannot be rejected. The results mean that there existed both long-run and short-run market integrations between Osogbo and the other selected markets. Thus, changes in ~~food-grain the price of maize~~ in rural markets would cause ~~food-grain the price of maize~~ in urban markets to adjust immediately.

~~Table 14: Wald Test for local rice market~~

~~Null hypothesis~~	~~F-statistics~~	~~p-value~~
~~Long-run market integration~~	~~0.20~~	~~0.85~~
~~Short-run market integration~~	~~0.19~~	~~0.75~~

Table 15 9: Wald Test for white Maize Market

Null hypothesis	F-statistics	Pp -value
-----------------	--------------	-----------

Long-run market integration	0.25	0.96
Short-run market integration	0.57	0.62

CONCLUSION

This study is an evaluation of market integration and price variation in marketing of ~~local rice and~~ white maize in Osun State, Nigeria. The trend analysis showed that the prices of ~~feed-grain~~ maize in the markets studied, moved in an upward trend every year. This is due to the fact that prices were higher in one year compared to other years. Results of Pearson correlation coefficient indicated that the rural and urban market price series for ~~both product types~~ maize were positively and significantly correlated. The stationary test indicated that the prices were not stationary at level form. The results of the maximum likelihood test showed that there were co-integrating vectors, which suggested that ~~rice and~~ maize markets were co-integrated and had short-run and long-run relationships. The result of the Granger causality test confirmed ~~that rural grain maize~~ markets ~~were~~ occupying the leadership position in price formation and transmission.

Recommendations

Based on the results of the study, the following are recommended:

- There is the need for efficient transmission of price information among the operators in the urban and rural markets through the establishment of market information centers to facilitate adequate communication and flow of information between markets.
- Accurate and timely information on ~~feed-grain the conditions of maize~~ will enable quicker response to market shocks, and market channel members could then be empowered to efficiently and effectively distribute ~~feed-grain maize~~ from surplus to deficit markets.
- Rural areas which have been shown in this study to be the market leaders should be the target of government developmental reforms. Incomes of rural people can be greatly enhanced with such incentives by government to intensify their production and marketing of ~~grain maize~~ which will

create greater opportunities for economic growth and development and eventually leads to market efficiency, increase technical and allocative efficiency of ~~grains~~-maize producers.

ACKNOWLEDGEMENTS

The authors would like to acknowledge the contributions of **EMOKARO, C.O** to this study. Gratitude also goes to the authorities of Osun State Agricultural Development Programme for the support rendered in the course of carrying out this study..

. Competing interests

No competing interests exist.

AUTHORS' CONTRIBUTIONS

. **DR. (MRS). G.O. ALUFOHAI** designed the study, wrote the protocol, wrote the first draft of the manuscript, managed the analyses of the study and approved the final manuscript. Ayantoyinbo A.A. performed the statistical analysis and carried out the field work. **Both authors managed the literature search.**

CONSENT (WHERE EVER APPLICABLE)

Not applicable.

ETHICAL APPROVAL (WHERE EVER APPLICABLE)

Not applicable.

REFERENCES

1. Obayagbona J, Eboiyehi O. Impact of agricultural output on per capita income in Nigeria: 1980-2011. Nigerian Journal of Agricultural and Development Economics. 2013;3(1):141-150.
2. World Bank. Development report, Washington D.C.; 2005
3. ~~NEAERLS (1996)~~ Prospects and problems of the 1996 Cropping season. A report of a study conducted by the National Agricultural Extension Research and Liaison Services; 1996.
4. Kolawale O, Ojo SO. Economic efficiency of small scale food crop production in Nigeria. J. For Soc. Sci., 2007;14(2):123-130.

5. Onu JI, Iliyasu, HA. An Economic Analysis of the Food Grain Market in Adamawa State, Nigeria. *World Journal of Agricultural Sciences* 2008;4(5):617-622,
6. Adetunji MO, Adesiyun IO. Economic Analysis of Plantain Marketing in Akinyele Local Government Area of Oyo State, Nigeria. *Int. Journal of Agricultural Economics and Rural Devevelopment*, 2008;1(1):15-21.
7. Olukosi JO, Isitor SU, Ode MO. *Introduction to Agricultural Marketing and Prices: Principles and Applications*. Third Edition, ~~G.U. publishers GU Abuja, Nigeria; 2007. Abuja Nigeria. Pp. 47-57.~~
8. National Population Commission (2006). Population Census, 2006. Federal Republic of Nigeria. [http://www.nigerianstat.gov.ng/national_bureau_of_statistics/official_gazatte_\(fgp_71/52007/2,500\(124\):pdf](http://www.nigerianstat.gov.ng/national_bureau_of_statistics/official_gazatte_(fgp_71/52007/2,500(124):pdf). Retrieved June,2012.
9. Osun State Official website. Osun State. <http://www.e-nigeria.net/osun.html>. retrieved on 03/10/2013.
10. Ogunbodede BA, Olakojo SA. "Development of Striga Ash Tolerant Hybrid Maize (Zea Mays L.) Varieties". *Tropical Agric Research and Extension* 2001;4(1):42-54.
11. Mohammad S. Supply response analysis of major crops in different agro-ecological zones in Punjab. *Pakistan Journal of Agricultural Research*, 2005;124(54):144-156.
12. Enders W. *Applied Econometric Time Series*, 3rd edition, New York: Wiley, 2010.
13. Widarjono A. *Econometrics: Theory and applications to economics and business*. Ekinosia Faculty of Economics, UII, Yogyakarta; 2007.
15. Dickey DA, Fuller WA. Distribution of the Estimators for Autoregressive Time Series with Unit Roots. *Journal of the American Statistical Association*, 1979;74(366): 427-431.
16. Dickey DA, Fuller WA. Likelihood Ratio Statistics for Autoregressive Time Series with a Unit Root. *Econometrica*, 1981;49(4):1057-1072.
17. Said SE, Dickey DA. Testing for unit roots in autoregressive-moving average models of unknown order. *Biometrika*, 1984;71:599-607.
18. Greene WH. *Econometric Analysis*. Second Edition. NewYork: Macmillan, 1993.
19. Engle RF, Granger CWJ. "Co-integration and Error Correction: Representation, Estimation and Testing, *Econometrica*, 1987;55(2):251-276.
20. Johansen S, Juselius K. Maximum Likelihood Estimation and Inference on Cointegration with Application to the Demand for Money, *Oxford Bulletin of Economics and Statistics*, 1990;52(2):169-210.
21. Rapsomanikis G, Hallam D, Conforti P. Market integration and price transmission in selected food and cash crop markets of developing countries: Review and Applications. 32005;225-229.
22. Oladapo MO, Momoh S. Food price differences and market integration in Oyo State, Nigeria. *International Journal of Agricultural Research*. 2007;2(1):69-74.
23. Ojiako IA, Ezedinma C, Asumugha GN, Nkang NM. Spatial Integration and Price Transmission in Selected Cassava Products' Markets in Nigeria: A Case of Lafun *World Applied Sciences Journal* 2012;18(9):1209-1219.

24. Hopcraft P. Grain marketing policies and institutions in Africa. Finance and development, 1987;24(1):37-43.
25. Ahmed R, Rustagi N. Marketing and price incentives in African and Asian Countries. In: Elz D, editor. Agricultural marketing strategy and pricing policy. International Bank for Reconstruction and Development, Washington, USA;1987.
26. Okoh RN, Egbon PC. The integration of Nigeria's rural and urban foodstuff markets. AERC Research Paper 151, *African Economic Research Consortium*, Nairobi, November 2005.
27. Mafimisebi TE. Long-Run price integration in the Nigerian Fresh fish market: Implication for development, Delhi Business Review, 2008;XI 9(1): 13-18.
29. Desi A, Yulius J. Integration of Rice Market Inter-Provinces of Rice Production Center in Indonesia. International Conference on Environment, Energy and Biotechnology IPCBEE 2012;33:42-51.
14. Chirwa EW. Food Marketing Reforms and Integration of Maize and Rice Markets in Malawi. Working Paper, School of Economics, University of East Anglia, 2000.

~~International Rice Research Institute (IRRI), (2001). Rice Statistics.~~

IRRI. World Rice Statistics. International Rice Research Institute, Philippines; 2001.

~~National Cereals Research Institute (NCRI) 2008. Bedeggi, Nigeria~~ NCRI. Metrological Station report. National Cereals Research Institute. Bedeggi, Nigeria; 2008 Food and Agriculture Organization. World Agriculture: towards 2010. An FAO Study. Alexandros N, editor. <http://www.fao.org/docrep/V4200E00.HTM> Retrieved on 1/10/2013

1. Negassa A, Myers R. estimating policy effects in special market efficiency: an extension to the parity bonds model. American Journal of Agricultural Economics 2007;89:338-352.

DEFINITIONS, ACRONYMS, ABBREVIATIONS

Central Bank of Nigeria (CBN), Food and Agriculture Organization (FAO), National Agricultural Extension Research and Liaison Services (NAERLS), National Population Commission (NPC).